

DEPARTMENT OF ECONOMICS, MANAGEMENT AND  
QUANTITATIVE METHODS

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B-74-3-B Time Series Econometrics

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Discussion of Exercise Sheet 2

1.

Autocorrelation structure of two AR(2). We discuss parts i. and ii. at the same time.

$$Y_t = c + \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t \text{ where } \varepsilon_t \text{ wn } (0, \sigma^2).$$

Discuss stationarity first.

When  $\phi_1 = 0.8$ ,  $\phi_2 = -0.8$ , the characteristic equation is

$(1 - 0.8z + 0.8z^2) = 0$ , Solution is:  $0.5 + 1.0i$ ,  $0.5 - 1.0i$  and  $|0.5 \pm 1| > 1$  so the process is stationary because both the roots are outside the unit circle. From the fact that it has complex roots we can also see it has a cycle.

When  $\phi_1 = -0.5$ ,  $\phi_2 = 0.3$ , the characteristic equation is

$(1 + 0.5z - 0.3z^2) = 0$ , Solution is:  $2.8403$ ,  $-1.1736$ , so the process is stationary because both the roots are outside the unit circle.

i. We are interested in the autocorrelation: for stationary processes, these are

$$\rho_j = \frac{\gamma_j}{\gamma_0}$$

where

$$\mu = E(Y_t) \text{ and } \gamma_j = E[(Y_t - \mu)(Y_{t-j} - \mu)]$$

First,

$$\begin{aligned} E(Y_t) &= E(c + \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t) \\ &= E(c) + \phi_1 E(Y_{t-1}) + \phi_2 E(Y_{t-2}) + E(\varepsilon_t) \end{aligned}$$

and, using stationarity,  $E(Y_t) = E(Y_{t-1}) = E(Y_{t-2}) = \mu$ ; moreover, because  $\varepsilon_t$  is white noise,  $E(\varepsilon_t) = 0$ , so

$$\mu = \frac{c}{1 - \phi_1 - \phi_2}.$$

Rewriting  $c = \mu(1 - \phi_1 - \phi_2)$ , our model becomes

$$(Y_t - \mu) = \phi_1(Y_{t-1} - \mu) + \phi_2(Y_{t-2} - \mu) + \varepsilon_t \text{ where } \varepsilon_t \text{ wn}(0, \sigma^2)$$

Of course, we could have skipped all this part if  $c = 0$ , in which case  $\mu = 0$  and we have directly

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t \text{ where } \varepsilon_t \text{ wn}(0, \sigma^2), \text{ and } \gamma_j = E(Y_t Y_{t-j})$$

So,

$$\begin{aligned} \gamma_{j \geq 1} &= E(Y_t \times Y_{t-j}) = E((\phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t) Y_{t-j}) \\ &= E(\phi_1 Y_{t-1} Y_{t-j}) + E(\phi_2 Y_{t-2} Y_{t-j}) + E(\varepsilon_t Y_{t-j}) \\ &= \phi_1 \gamma_{j-1} + \phi_2 \gamma_{j-2} \end{aligned}$$

where we used

$$E(Y_{t-1} Y_{t-j}) = E(Y_t Y_{t-(j-1)}) = \gamma_{j-1}, \quad E(Y_{t-2} Y_{t-j}) = E(Y_t Y_{t-(j-2)}) = \gamma_{j-2}$$

because of stationarity, and

$$E(\varepsilon_t Y_{t-j}) = 0 \text{ for } j \geq 1$$

because  $\varepsilon_t$  is white noise so it does not depend on the past ( $E(\varepsilon_t \varepsilon_{t-j}) = 0$  for  $j \geq 1$ ), while  $Y_{t-j}$  is a past value (when  $j \geq 1$ ). So

$$\gamma_j = \phi_1 \gamma_{j-1} + \phi_2 \gamma_{j-2} \text{ for } j \geq 1$$

(and notice that  $\gamma_j = \gamma_{-j}$ , so  $\gamma_1 = \phi_1 \gamma_0 + \phi_2 \gamma_1$ ). Dividing by  $\gamma_0$ ,

$$\rho_{j \geq 1} = \phi_1 \rho_{j-1} + \phi_2 \rho_{j-2},$$

(Yule Walker equations) which we initialise by setting

$$\rho_1 = \phi_1 \rho_0 + \phi_2 \rho_{-1}$$

(again using stationarity,  $\rho_1 = \rho_{-1}$ ) so

$$\rho_1 = \frac{\phi_1}{1 - \phi_2}$$

and for  $\rho_2$  just notice that  $\rho_0 = 1$ ,

$$\rho_2 = \phi_1 \rho_1 + \phi_2$$

and iterating.

Note. It is worth mentioning here that we did not need to compute  $\gamma_0$  to derive these autocorrelations.

The plot is very different, a cycle can be observed for the process having complex roots in the characteristic equation, while for the other process the autocorrelations change sign at every step.

	$\rho_1$	$\rho_2$	$\rho_3$	$\rho_4$	$\rho_5$
<i>if</i> $\phi_1 = 0.8, \phi_2 = -0.8$	0.444	-0.444	-0.711	-0.213	0.398
<i>if</i> $\phi_1 = -0.5, \phi_2 = 0.3$	-0.714	0.657	-0.542	0.468	-0.397

iii IRF - A plot of  $\frac{\partial Y_t}{\partial \varepsilon_{t-j}}$  (against  $j$ ) is called Impulse Response Function. For a process  $Y_t$  that admits

$$Y_t = \mu + \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j}$$

for  $\varepsilon_t$  such that, for any  $t$ ,

$$\begin{aligned} E(\varepsilon_t) &= 0, E(\varepsilon_t^2) = \sigma^2, \\ E(\varepsilon_t \varepsilon_\tau) &= 0 \text{ if } \tau \neq t \end{aligned}$$

notice that

$$\frac{\partial Y_t}{\partial \varepsilon_{t-j}} = \psi_j$$

so  $\psi_j$  is the effect on  $Y_t$  of a shock that took place  $t - j$  periods before.

It may also be of interest to compute the  $\psi_j$  in the IRF (Wold decomposition):

Any ARMA(p,q) can be represented as  $\phi(L)Y_t = \theta(L)\varepsilon_t$ , where  $\phi(L) = 1 - \phi_1L - \phi_2L^2 - \dots - \phi_pL^p$ ,  $\theta(L) = 1 + \theta_1L + \dots + \theta_qL^q$  and stationarity ensures  $Y_t = \phi^{-1}(L)\theta(L)\varepsilon_t$ .

We are looking for the parameters  $\psi_j$  in the infinite polynomial

$\psi(L) = 1 + \psi_1L + \dots$ , such that  $Y_t = \psi(L)\varepsilon_t$ :

this means that  $\phi^{-1}(L)\theta(L) = \psi(L)$ , and then  $\theta(L) = \phi(L)\psi(L)$

this is

$$\begin{aligned} & 1 + \theta_1L + \theta_2L^2 + \theta_3L^3 + \dots + \theta_qL^q \\ = & (1 - \phi_1L - \phi_2L^2 - \phi_3L^3 - \dots - \phi_pL^p) (1 + \psi_1L + \psi_2L^2 + \psi_3L^3 + \dots) \end{aligned}$$

$$\begin{aligned} & 1 + \theta_1L + \theta_2L^2 + \theta_3L^3 \dots + \theta_qL^q \\ = & 1 - \phi_1L + \psi_1L - \psi_1\phi_1L^2 + \psi_2L^2 - \phi_2L^2 - \phi_3L^3 - \phi_2\psi_1L^3 - \phi_1\psi_2L^3 + \psi_3L^3 + \dots \end{aligned}$$

since this is an identity the elements of the same order must be equal,

so,

for the terms of order  $L$ ,  $\theta_1 = \psi_1 - \phi_1$ , which means  $\psi_1 = \theta_1 + \phi_1$ ,

for the terms of order  $L^2$ ,  $\theta_2 = \psi_2 - \phi_2 - \psi_1\phi_1$ , which means  $\psi_2 = \theta_2 + \phi_2 + \psi_1\phi_1$  (notice that  $\psi_1$  is known at this point, because it was determined in the previous step)

for the terms of order  $L^3$ ,  $\theta_3 = \psi_3 - \phi_3 - \psi_1\phi_2 - \phi_2\psi_1$ , which means  $\psi_3 = \theta_3 + \phi_3 + \psi_1\phi_2 + \phi_2\psi_1$

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In this case  $\theta_{j \geq 1} = 0$ , so we have

$$1 = 1 - \phi_1L + \psi_1L - \psi_1\phi_1L^2 + \psi_2L^2 - \phi_2L^2 - \phi_3L^3 - \phi_2\psi_1L^3 - \phi_1\psi_2L^3 + \psi_3L^3 \dots$$

and then

$$\psi_1 = \phi_1,$$

$$\psi_2 = \phi_2 + \psi_1\phi_1$$

$$\psi_3 = \phi_1\psi_2 + \phi_2\psi_1$$

i.e.

$$\psi_{j \geq 1} = \phi_2\psi_{j-2} + \phi_1\psi_{j-1} \text{ (recall } \psi_0 = 1).$$

For the given parameters, these weights are

When  $\phi_1 = 0.8$ ,  $\phi_2 = -0.8$ :

$$\psi_1 = 0.8, \psi_2 = -0.16, \psi_3 = -0.768, \psi_4 = -0.4864, \psi_5 = 0.2253$$

(notice again the cyclical component);

When  $\phi_1 = -0.5$ ,  $\phi_2 = 0.3$ :

$$\psi_1 = -0.5, \psi_2 = 0.55, \psi_3 = -0.425, \psi_4 = 0.3775, \psi_5 = -0.31625.$$

**2.**

i. The process is stationary ( $|-0.5| < 1$ ) and invertible ( $|0.7| < 1$ ).

ii. Rewriting the ARMA(p,q) as  $\phi(L)Y_t = \theta(L)\varepsilon_t$ , where

$$\phi(L) = 1 - \phi_1L - \phi_2L^2 - \dots - \phi_pL^p, \theta(L) = 1 + \theta_1L + \dots + \theta_qL^q$$

and the polynomial of the MA( $\infty$ ) given by the Wold decomposition as

$$Y_t = \psi(L)\varepsilon_t \text{ where } \psi(L) = 1 + \psi_1L + \dots,$$

$$\text{then } \phi^{-1}(L)\theta(L) = \psi(L), \text{ and then } \theta(L) = \phi(L)\psi(L)$$

$$\text{this is } 1 + \theta_1L + \dots + \theta_qL^q = (1 - \phi_1L - \phi_2L^2 - \dots - \phi_pL^p)(1 + \psi_1L + \dots)$$

$$1 + \theta_1L + \dots + \theta_qL^q = 1 - \phi_1L + \psi_1L - \psi_1\phi_1L^2 + \psi_2L^2 - \phi_2L^2 \dots$$

since this is an identity the elements of the same order must be equal,

so,

$$\text{for the terms of order } L, \theta_1 = \psi_1 - \phi_1, \text{ which means } \psi_1 = \theta_1 + \phi_1,$$

for the terms of order  $L^2$ ,  $\theta_2 = \psi_2 - \phi_2 - \psi_1\phi_1$ , which mean  $\psi_2 = \theta_2 + \phi_2 + \psi_1\phi_1$  (notice that  $\psi_1$  is known at this point, because it was determined in the previous step)

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In this case

$$1 + \theta L = 1 - \phi L + \psi_1L - \psi_1\phi L^2 + \psi_2L^2 - \phi\psi_2L^3 + \psi_3L^3 \dots$$

and then

$$\psi_1 = \phi + \theta = 1.2$$

$$\psi_2 = \phi\psi_1 = 0.6$$

$$\psi_3 = \phi\psi_2 = 0.3$$

and, in general,

$$\psi_{j \geq 2} = \phi\psi_{j-1}.$$

You can also prove it by looking at  $Y_t = \phi Y_{t-1} + \xi_t$  where  $\xi_t = \varepsilon_t + \theta\varepsilon_{t-1}$ .

Then,

$$\begin{aligned} Y_t &= \sum_{j=0}^{\infty} \phi^j \xi_{t-j} = \sum_{j=0}^{\infty} \phi^j (\varepsilon_{t-j} + \theta\varepsilon_{t-j-1}) = \sum_{j=0}^{\infty} \phi^j \varepsilon_{t-j} + \theta \sum_{j=0}^{\infty} \phi^j \varepsilon_{t-j-1} \\ &= \sum_{j=0}^{\infty} \phi^j \varepsilon_{t-j} + \theta \sum_{l=1}^{\infty} \phi^{l-1} \varepsilon_{t-l} = \varepsilon_t + \sum_{j=1}^{\infty} \phi \phi^{j-1} \varepsilon_{t-j} + \theta \sum_{j=1}^{\infty} \phi^{j-1} \varepsilon_{t-j} \\ &= \varepsilon_t + (\theta + \phi) \sum_{j=1}^{\infty} \phi^{j-1} \varepsilon_{t-j} \end{aligned}$$

iii. Before computing the autocorrelations, notice that

$$Cov(Y_t, \varepsilon_t) = Cov(\phi Y_{t-1} + \varepsilon_t + \theta\varepsilon_{t-1}, \varepsilon_t) =$$

$$= Cov(\phi Y_{t-1}, \varepsilon_t) + Cov(\varepsilon_t, \varepsilon_t) + Cov(\theta\varepsilon_{t-1}, \varepsilon_t) = 0 + \sigma^2 + 0.$$

Next,

$\gamma_0 = \text{Var}(Y_t) = \text{Var}(\phi Y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1}) =$   
 $= \text{Var}(\phi Y_{t-1}) + \text{Var}(\varepsilon_t) + \text{Var}(\theta \varepsilon_{t-1}) + 2\text{Cov}(\phi Y_{t-1}, \theta \varepsilon_{t-1})$  (the other covariances are 0) so

$$\begin{aligned} \gamma_0 &= \phi^2 \text{Var}(Y_{t-1}) + \text{Var}(\varepsilon_t) + \theta^2 \text{Var}(\varepsilon_{t-1}) + 2\phi\theta \text{Cov}(Y_{t-1}, \varepsilon_{t-1}) = \\ \gamma_0 &= \phi^2 \gamma_0 + \sigma^2 + \theta^2 \sigma^2 + 2\phi\theta \sigma^2 \text{ using stationarity,} \\ \gamma_0 &= \frac{1+\theta^2+2\phi\theta}{1-\phi^2} \sigma^2 \end{aligned}$$

and

$$\begin{aligned} \gamma_1 &= \text{Cov}(Y_t, Y_{t-1}) = \text{Cov}(\phi Y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1}, Y_{t-1}) = \phi \gamma_0 + 0 + \theta \sigma^2 \\ \text{so } \gamma_1 &= \left( \frac{1+\theta^2+2\phi\theta}{1-\phi^2} \phi + \theta \right) \sigma^2 = \frac{\phi + \phi\theta^2 + 2\phi^2\theta + \theta - \phi^2\theta}{1-\phi^2} \sigma^2 = \frac{\phi + \phi\theta^2 + \phi^2\theta + \theta}{1-\phi^2} \sigma^2 = \\ &= \frac{\phi + \theta + \phi\theta(\theta + \phi)}{1-\phi^2} \sigma^2 = \frac{(\theta + \phi)(1 + \phi\theta)}{1-\phi^2} \sigma^2 \end{aligned}$$

and

$$\gamma_2 = \text{Cov}(Y_t, Y_{t-2}) = \text{Cov}(\phi Y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1}, Y_{t-2}) = \text{Cov}(\phi Y_{t-1}, Y_{t-2}) = \phi \gamma_1,$$

and in general

$$\gamma_{j \geq 2} = \text{Cov}(Y_t, Y_{t-j \geq 2}) = \text{Cov}(\phi Y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1}, Y_{t-j \geq 2}) = \text{Cov}(\phi Y_{t-1}, Y_{t-j \geq 2}) = \phi \gamma_{j-1},$$

So

$$\rho_1 = \frac{(\theta + \phi)(1 + \phi\theta)}{1 + \theta^2 + 2\phi\theta}$$

$$\rho_{j \geq 2} = \phi \rho_j$$

so the autocorrelation function has a bump at the first lag, but behaves like an AR(1) otherwise (notice that this argument could be generalised in order to recognise any ARMA(p,q) model).

So,

$$\rho_1 = \frac{(\theta + \phi)(1 + \phi\theta)}{1 + \theta^2 + 2\phi\theta} = \frac{(0.7 + 0.5)(1 + 0.5 * 0.7)}{1 + 0.7^2 + 2 * 0.5 * 0.7} = 0.73973$$

$$\rho_2 = 0.5 * \rho_1 = 0.36987$$

$$\rho_3 = 0.5 * \rho_2 = 0.5 * 0.36987 = 0.18494$$

An alternative way to compute  $\rho_1$  is to use the  $MA(\infty)$  representation:

$$\begin{aligned} \text{then } \gamma_0 &= \sum_{k=0}^{\infty} \psi_k^2 \sigma^2 \text{ and } \sum_{k=0}^{\infty} \psi_k^2 = 1^2 + (\phi + \theta)^2 \sum_{k=1}^{\infty} \phi^{(k-1)2} = 1^2 + \\ &(\phi + \theta)^2 \sum_{l=0}^{\infty} \phi^{2l} = 1^2 + \frac{(\phi + \theta)^2}{1 - \phi^2} = \\ &1^2 + \frac{(0.5 + 0.7)^2}{1 - 0.5^2} = 2.92 \end{aligned}$$

$$\text{and } \gamma_j = \sum_{k=0}^{\infty} \psi_k \psi_{k+j} \sigma^2 \text{ and, for } j = 1, \sum_{k=0}^{\infty} \psi_k \psi_{k+1} = 1 * (\theta + \phi) + (\theta + \phi)^2 \sum_{k=1}^{\infty} \phi^{(k-1) + (k-1+1)} =$$

$$1 * (\theta + \phi) + (\theta + \phi)^2 \phi \sum_{k=1}^{\infty} \phi^{2(k-1)} = (\theta + \phi) + \frac{(\theta + \phi)^2 \phi}{1 - \phi^2} = 0.7 + 0.5 + 0.5 \frac{(0.7 + 0.5)^2}{1 - 0.5^2} = 2.16$$

$$\rho_1 = \frac{2.16 \sigma^2}{2.92 \sigma^2} = 0.73973$$

(could use this procedure for higher lags as well)

**3.**

Just factorise  $Y_t = 0.7Y_{t-1} - 0.1Y_{t-2} + \varepsilon_t + 0.5\varepsilon_{t-1} - 0.14\varepsilon_{t-2}$

$$(1 - 0.7L + 0.1L^2) Y_t = (1 + 0.5L - 0.14L^2) \varepsilon_t,$$

$$[(1 - 0.5L)(1 - 0.2L)] Y_t = [(1 + 0.7L)(1 - 0.2L)] \varepsilon_t,$$

(verify that the model is stationary, then, because  $|0.5| < 1$  and  $|0.2| < 1$ )

so the factor  $(1 - 0.2L)$  is common, and the model can be reparametrised

as

$$(1 - 0.5L) Y_t = (1 + 0.7L) \varepsilon_t,$$

$$Y_t = 0.5Y_{t-1} + \varepsilon_t + 0.7\varepsilon_{t-1}$$

for which we already computed the autocorrelation function.

**4.**

i. Assuming that  $E(Y_t) = 0$ ,

$$\widehat{Y}_{t+1|t,t-1,\dots,1} = \alpha_1^{(t)} Y_t + \alpha_2^{(t)} Y_{t-1} + \dots + \alpha_t^{(t)} Y_1$$

where, letting  $\gamma_j = E(Y_t Y_{t+j})$ ,

$$\begin{pmatrix} \alpha_1^{(t)} \\ \alpha_2^{(t)} \\ \dots \\ \alpha_t^{(t)} \end{pmatrix} = \begin{pmatrix} \gamma_0 & \gamma_1 & \dots & \gamma_{t-1} \\ \gamma_1 & \gamma_0 & \dots & \gamma_{t-2} \\ \dots & \dots & \dots & \dots \\ \gamma_{t-1} & \gamma_{t-2} & \dots & \gamma_0 \end{pmatrix}^{-1} \begin{pmatrix} \gamma_1 \\ \gamma_2 \\ \dots \\ \gamma_t \end{pmatrix}$$

ii. Inverting this matrix is computationally intensive, when  $t$  is large. As an alternative, setting  $\widehat{\varepsilon}_1 = 0$ , we may compute

$$\widehat{\varepsilon}_2 = Y_2 - 0.2Y_1$$

$$\widehat{\varepsilon}_s = Y_s - 0.2Y_{s-1} - 0.6\widehat{\varepsilon}_{s-1} \text{ for } s \geq 1$$

and finally,

$$\widehat{Y}_{t+1|t,t-1,\dots,1,\widehat{\varepsilon}_1=0} = 0.2Y_t + 0.6\widehat{\varepsilon}_t$$